

Bayesian inference for expensive computer models in chemical engineering

Peter L. W. Man and Markus Kraft

CoMo Group
Department of Chemical Engineering and Biotechnology
University of Cambridge

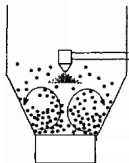
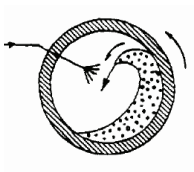
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Granulation

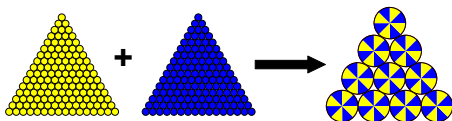
■ Products



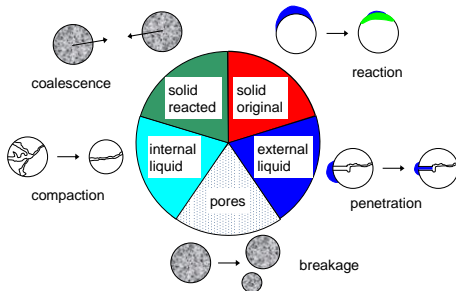
■ Devices



- Lump small particles into bigger entities
 - Improve **handling** (storage, transport, safety,...)
 - Creation of **micro mixtures** (segregation, distribution of active components)
- Desire knowledge about:
 - **Influences** of precursors, process design and conditions **on final product**



Granulation model



transformation	unknown parameter
coalescence	k_{coag}
compaction	k_{compact}
breakage	k_{break}
penetration	k_{pen}
reaction	k_{reac}

Aim: Regress for some data points \mathbf{y}_{obs} over input space \mathcal{Z}

- Assume

$$y_{\text{obs}}(z) = y(z) + \epsilon \quad \text{where} \quad \epsilon \sim \mathcal{N}(0, \tau^2) \quad (1)$$

- Underlying $y(z)$ is *unknown*, so endow with prior distribution :
 - $y(\cdot) \sim \text{GP}(\mu(\cdot), \sigma^2 \Sigma(\cdot, \cdot))$
 - mean function $\mu(z)$
 - correlation function $\Sigma(z_1, z_2)$

- Make observations $\mathbf{y}_{\text{obs}} := y(\mathbf{z}) + \epsilon$ at $\mathbf{z} = (z_1, \dots, z_n)^\top$
- Want predictions at $\tilde{\mathbf{z}}$

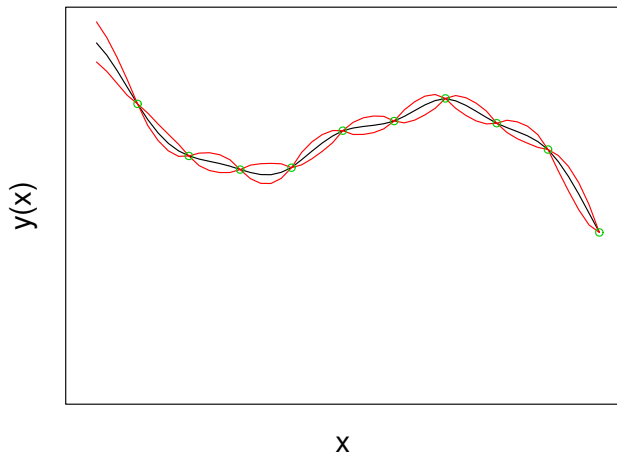
Posterior of $y(\tilde{\mathbf{z}}) | \mathbf{y}_{\text{obs}}$:

- $y(\tilde{\mathbf{z}}) | \mathbf{y}_{\text{obs}}$ has a multivariate normal distribution with

$$\begin{aligned} \mathbb{E} [y(\tilde{\mathbf{z}}) | \mathbf{y}_{\text{obs}}] \\ = \mu(\tilde{\mathbf{z}}) + \sigma^2 \Sigma(\tilde{\mathbf{z}}, \mathbf{z}) [\sigma^2 \Sigma(\mathbf{z}, \mathbf{z}) + \tau^2 I]^{-1} (\mathbf{y}_{\text{obs}} - \mu(\mathbf{z})) \quad (2) \end{aligned}$$

$$\begin{aligned} \text{Var} (y(\tilde{\mathbf{z}}) | \mathbf{y}_{\text{obs}}) \\ = \sigma^2 \Sigma(\tilde{\mathbf{z}}, \tilde{\mathbf{z}}) - \sigma^4 \Sigma(\tilde{\mathbf{z}}, \mathbf{z}) [\sigma^2 \Sigma(\mathbf{z}, \mathbf{z}) + \tau^2 I]^{-1} \Sigma(\tilde{\mathbf{z}}, \mathbf{z})^\top \quad (3) \end{aligned}$$

Regression IV



Choices of $\Sigma(z_1, z_2)$:

- Exponential :

$$\Sigma(z_1, z_2) = \prod_{k=1}^{\dim(\mathcal{Z})} \exp\left(-\phi_k \cdot |z_{1k} - z_{2k}|^{\phi_0}\right) \quad (4)$$

- Matérn :

$$\Sigma(z_1, z_2) = \frac{1}{2^{\nu-1}\Gamma(\nu)} \left(\frac{2\nu^{\frac{1}{2}}\|z_1 - z_2\|}{\rho}\right)^{\nu} \mathcal{K}_{\nu}\left(\frac{2\nu^{\frac{1}{2}}\|z_1 - z_2\|}{\rho}\right) \quad (5)$$

where \mathcal{K} is modified Bessel function of 2nd kind.

- \mathcal{X} is space of *process conditions* or covariates
- Θ is space of unknown parameters of a model
- $\mathcal{Z} := \mathcal{X} \times \Theta$

- Underlying **Simulator** function $\eta_s(z)$:
 - $\eta_s(z) \sim \text{GP}(\mu_s(\cdot), \sigma_s^2 \Sigma_s(\cdot, \cdot))$
- Underlying **Experimental** function $\eta_e(x)$:
 - Fix calibrated value of $\theta = \theta_c$ (unknown)
 - $\eta_e(x) := \eta_s(x, \theta_c) + \delta(x)$
 - $\delta(x) \sim \text{GP}(\mu_e(\cdot), \sigma_e^2 \Sigma_e(\cdot, \cdot))$
- $\implies \eta_e \sim \text{GP}(\mu_s(\cdot, \cdot) + \mu_e(\cdot), \sigma_s^2 \Sigma_s(\cdot, \cdot) + \sigma_e^2 \Sigma_e(\cdot, \cdot))$

- Observed **Experimental** data :
 - Have $\mathbf{y}_e := (\eta_e(x_1), \dots, \eta_e(x_{n_e}))^\top + \boldsymbol{\epsilon}_e$
- Observed **Simulator** data :
 - Have $\mathbf{y}_s := (\eta_s(x_1^*, \theta_1^*), \dots, \eta_s(x_{n_e}^*, \theta_{n_e}^*))^\top + \boldsymbol{\epsilon}_s$
- Full data $\mathbf{d} := (\mathbf{y}_s^\top, \mathbf{y}_e^\top)^\top$
- Observation errors $\boldsymbol{\epsilon}_e, \boldsymbol{\epsilon}_s$ for simplicity:
 - $\boldsymbol{\epsilon}_e \sim \mathcal{N}(0, \tau_e^2 I)$ and $\boldsymbol{\epsilon}_s \sim \mathcal{N}(0, \tau_s^2 I)$

- GP Priors
 - GP Matérn parameters $\phi_e, \phi_s \sim U[0, \phi_{\text{upper}}]$
 - GP mean linear parameters β_e, β_s —constant prior
 - GP variance parameters σ_e^2, σ_s^2 —inverse gamma
- Observation error priors
 - τ_e^2, τ_s^2 —inverse gamma
- Prior for θ is given by user
- All unknown values — represent as $(\theta, \beta_s, \beta_e, \xi)$

- Posterior for $(\theta, \beta_s, \beta_e, \xi)$:

$$\pi(\theta, \beta_s, \beta_e, \xi | \mathbf{d}) \propto \underbrace{f(\mathbf{d} | \theta, \beta_s, \beta_e, \xi)}_{\text{gaussian likelihood}} \times \underbrace{\pi(\theta)\pi(\beta_s)\pi(\beta_e)\pi(\xi)}_{\text{prior}} \quad (6)$$

- Integrate out β_s and β_e *analytically*
- Integrating out ξ is harder - do this via Wang-Laudau sampling
- Do similar for prediction of $\eta_e(x)$ and $\eta_s(x, \theta)$.

Example - simulator function $\eta_s(x, \theta)$

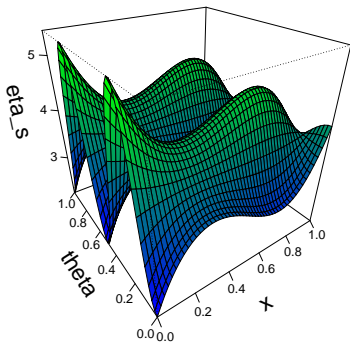
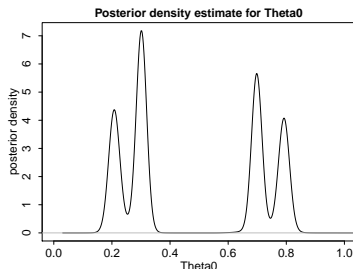
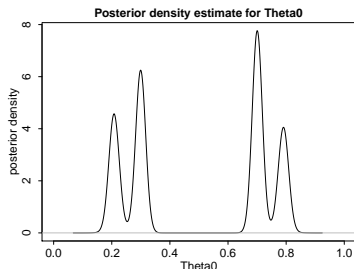


Figure: True $\eta_s(x, \theta)$

$$\mu_s(x, \theta) := \frac{1}{4} [6(x - 0.15)(x - 0.5)(x - 0.85) - 0.27] [10 \cos(4\pi\theta) + 0.3] + \text{const}$$
$$\theta_c = 0.2$$

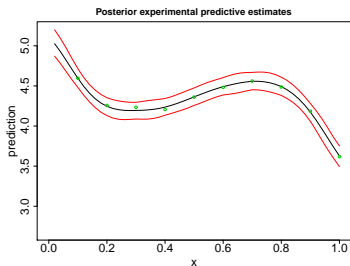
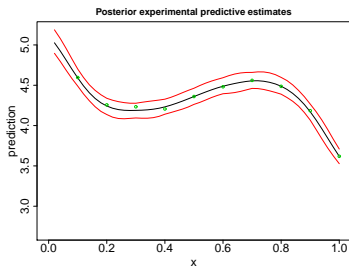
Example - θ -posterior



(a) assuming $\tau_s^2 = 0$, τ_e^2 unknown, $\delta(x) \equiv 0$
(b) assuming $\tau_s^2 = 0$, τ_e^2 and $\delta(x)$ unknown

Figure: θ posterior plots making different assumptions

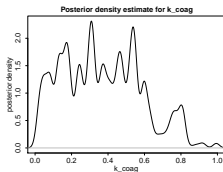
Example - $y_e(x)$ -posterior prediction + noise



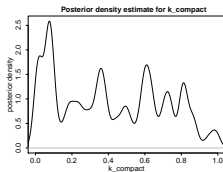
(a) assuming $\tau_s^2 = 0$, τ_e^2 unknown, (b) assuming $\tau_s^2 = 0$, τ_e^2 and $\delta(x)$ unknown
 $\delta(x) \equiv 0$ unknown

Figure: $y_e(x)$ (+noise) posterior plots making different assumptions

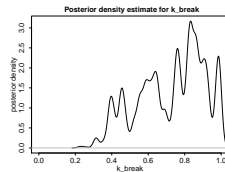
Granulation - (k_{coag} , k_{compact} , k_{break} , k_{pen} , k_{reac})-posteriors



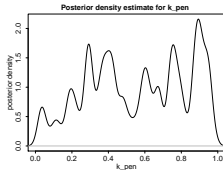
(a) k_{coag} posterior



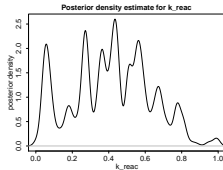
(b) k_{compact} posterior



(c) k_{break} posterior



(d) k_{pen} posterior



(e) k_{reac} posterior

Figure: Assuming $\tau_s^2 = 0$, $\tau_e^2, \delta(x)$ unknown