

# APTS - Survival Analysis

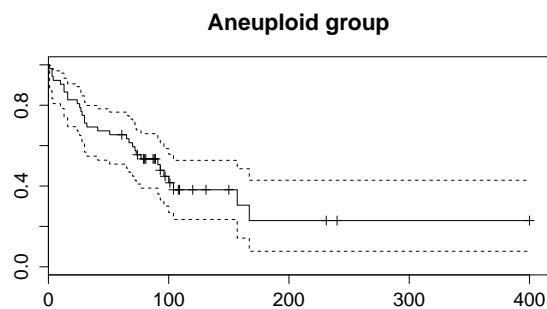
## Lab Session 1 - Solutions

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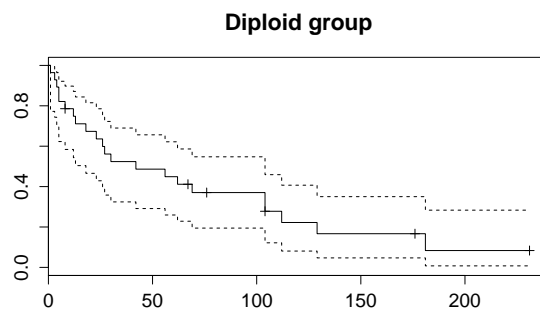
August 23, 2017

```
1. > install.packages("survival")
> library("survival")
> install.packages("KMsurv")
> library("KMsurv")
> data(tongue)
> tongue

> ane=subset(tongue,type==1)
> fit_ane=survfit(Surv(time,delta)~1,data=ane,conf.type="log-log")
> plot(fit_ane)
> title(main="Aneuploid group")
```



```
> dip=subset(tongue,type==2)
> fit_dip=survfit(Surv(time,delta)~1,data=dip,conf.type="log-log")
> plot(fit_dip)
> title(main="Diploid group")
```



```
> fit_ane
Call: survfit(formula = Surv(time, delta) ~ 1, data = ane, conf.type = "log-log")
```

n	events	median	0.95LCL	0.95UCL
52	31	93	65	157

```
> fit_dip
Call: survfit(formula = Surv(time, delta) ~ 1, data = dip, conf.type = "log-log")
```

n	events	median	0.95LCL	0.95UCL
28	22	42	18	104

2. > data(burn)

```
> burn
```

```
> survdiff(Surv(T3,D3)~Z1,data=burn)
```

```
Call:
```

```
survdiff(formula = Surv(T3, D3) ~ Z1, data = burn)
```

	N	Observed	Expected	(O-E) <sup>2</sup> /E	(O-E) <sup>2</sup> /V
Z1=0	70	28	21.4	2.07	3.79
Z1=1	84	20	26.6	1.66	3.79

Chisq= 3.8 on 1 degrees of freedom, p= 0.0515

```
> attach(burn)
```

```
> burn$area[Z4<=29] = 1
```

```
> burn$area[Z4>=30 & Z4<=50] = 2
```

```
> burn$area[Z4>=51] = 3
```

```
> survdiff(Surv(T3,D3)~Z1+strata(area),data=burn)
```

```
Call:
```

```
survdiff(formula = Surv(T3, D3) ~ Z1 + strata(area), data = burn)
```

	N	Observed	Expected	(O-E) <sup>2</sup> /E	(O-E) <sup>2</sup> /V
Z1=0	70	28	21.6	1.87	3.61
Z1=1	84	20	26.4	1.53	3.61

Chisq= 3.6 on 1 degrees of freedom, p= 0.0574

3. (a) (i) The number of distinct event times : 5

(ii) The ordered event times :

$j$	1	2	3	4	5
$y_{(j)}$	2	5	7	9	16

(iii) The size of the risk set at time  $y_{(1)}$  :  $y_{(1)} = 2$  and  $R_{(1)} = 10$  since all subjects are alive just before time  $t = 2$

(b) Calculation by hand (without using R) :

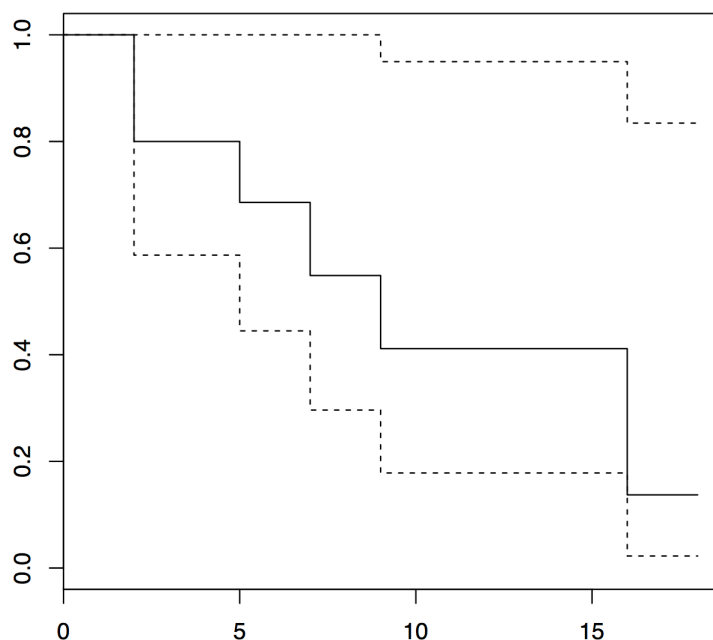
(i) Kaplan-Meier estimator :  $\hat{S}_{KM}(t) = \prod_{j:y_{(j)} \leq t} \left(1 - \frac{d_{(j)}}{R_{(j)}}\right)$

(ii) Nelson-Aalen estimator :  $\hat{S}_{NA}(t) = \prod_{j:y_{(j)} \leq t} \exp\left(-\frac{d_{(j)}}{R_{(j)}}\right)$

$j$	$y_{(j)}$	$d_{(j)}$	$R_{(j)}$	$1 - \frac{d_{(j)}}{R_{(j)}}$	$\hat{S}_{KM}(t)$	$\exp\left(-\frac{d_{(j)}}{R_{(j)}}\right)$	$\hat{S}_{NA}(t)$
1	2	2	10	0.800	0.800	0.819	0.819
2	5	1	7	0.857	0.686	0.867	0.71
3	7	1	5	0.800	0.549	0.819	0.581
4	9	1	4	0.750	0.411	0.779	0.523
5	16	2	3	0.333	0.137	0.513	0.232

(c) Graphical representation of the Kaplan-Meier estimator :

```
> install.packages("survival")
> y <- c(3,5,7,2,18,16,2,9,16,5)
> d <- c(0,1,1,1,0,1,1,1,1,0)
> library(survival)
> KM = survfit(Surv(y,d)~1)
> plot(KM)
```



(d) Comparison of the results :

```
> summary(KM)
```

```
Call: survfit(formula = Surv(y, d) ~ 1)
```

time	n.risk	n.event	survival	std.err	lower 95% CI	upper 95% CI
2	10	2	0.800	0.126	0.5868	1.000
5	7	1	0.686	0.151	0.4447	1.000
7	5	1	0.549	0.172	0.2963	1.000
9	4	1	0.411	0.176	0.1782	0.950
16	3	2	0.137	0.126	0.0225	0.834

4. (a) Log-rank test :

$$\begin{aligned}
 U &= \sum_{j=1}^r w(y_{(j)}) (O_j - E_j) \\
 &= \sum_{j=1}^r w(y_{(j)}) \left( d_{(j)1} - \frac{d_{(j)} R_{(j)1}}{R_{(j)}} \right)
 \end{aligned}$$

with  $\frac{U}{\sqrt{\text{Var}(U)}} \sim N(0, 1)$  and

$$\text{Var}(U) = \sum_{j=1}^r w^2(y_{(j)}) \frac{d_{(j)} \frac{R_{(j)1}}{R_{(j)}} \left( 1 - \frac{R_{(j)1}}{R_{(j)}} \right) (R_{(j)} - d_{(j)})}{R_{(j)} - 1}.$$

		group 1		group 2					Var(U)		
<i>j</i>	<i>y<sub>(j)</sub></i>	<i>d<sub>(j)1</sub></i>	<i>R<sub>(j)1</sub></i>	<i>d<sub>(j)2</sub></i>	<i>R<sub>(j)2</sub></i>	<i>d<sub>(j)</sub></i>	<i>R<sub>(j)</sub></i>	<i>E<sub>(j)</sub></i>	<i>N<sub>(j)</sub></i>	<i>D<sub>(j)</sub></i>	<i>N<sub>(j)</sub>/D<sub>(j)</sub></i>
1	4.1	1	6	0	6	1	12	0.5	2.75	11	0.25
2	9.7	0	4	1	6	1	10	0.4	2.16	9	0.24
3	10	2	4	1	5	3	9	1.333	4.44	8	0.555
4	17.2	1	1	0	2	1	3	0.333	0.44	2	0.22
5	19.7	0	0	1	2	1	2	0	0.00	1	0.00
<i>Total</i>		4		3		7		2.566			1.265

Hence,

$$U^{obs} = \frac{U}{\sqrt{Var(U)}} = \frac{1.434}{\sqrt{1.265}} = 1.275$$

We reject  $H_0$  if  $|U^{obs}| > z_{1-\alpha/2} = 1.96$ .

We have  $|U^{obs}| = 1.275 < z_{1-\alpha/2} = 1.96$ .

Hence, we do not reject  $H_0$ .

*P-value* =  $2 \times P(Z > 1.275) = 2 \times 0.101 = 0.202 > 0.05$

(b) Comparison of the results with the function `survdiff` :

```
> y = c(4.1,7.8,10,10,12.3,17.2,
+       9.7,10,11.1,13.1,19.7,24.1)
> d = c(1,0,1,1,0,1,
+       1,1,0,0,1,0)
> group = c( rep(1,6), rep(2,6) )
> cbind(y, d, group)
```

```
      y  d group
[1,] 4.1 1     1
[2,] 7.8 0     1
[3,] 10.0 1    1
[4,] 10.0 1    1
[5,] 12.3 0     1
[6,] 17.2 1     1
[7,]  9.7 1     2
[8,] 10.0 1     2
[9,] 11.1 0     2
[10,] 13.1 0    2
[11,] 19.7 1     2
[12,] 24.1 0     2
```

```
> library(survival)
> survdiff(Surv(y,d)~group)
```

Call:

```
survdiff(formula = Surv(y, d) ~ group)
```

	N	Observed	Expected	(O-E) <sup>2</sup> /E	(O-E) <sup>2</sup> /V
group=1	6	4	2.57	0.800	1.62
group=2	6	3	4.43	0.463	1.62

Chisq= 1.6 on 1 degrees of freedom, p= 0.203

The conclusion is the same.