Identifying Direct and Indirect effects

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Outline

Basics

- DAGs and intervention
- Causal?
- Examples
- Regimes
 - Heuristic
 - Formal
- Definition of effects
- Identification of effects

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- Extensions
- Conclusions

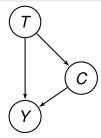
Basics

Conditional Independence

$$p(A, B|C) = p(A|C) \times p(B|C) \Rightarrow A \bot B|C$$

DAGs

- Directed Acyclic Graphs
- Encode conditional independence
- More than one can encode same CI like intuitve DAG for direct and indirect effects.



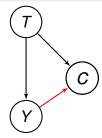
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Basics continued

Notation

- T treatment: 0,1
- **F**_T treatment assignment: $0, 1, \emptyset$
- C intermediate variable: 0,1
- Y outcome/response variable: y

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- T treatment: 0,1
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F_T : the regime indicator

•
$$p(T = t | X, F_T = t) = 1$$
 and

i.e. T is drawn from its "natural" or "observational" distribution

- F_T is an intervention variable, treatment assignment indicator, regime indicator.
- DAGs with intervention variables = Augmented DAGs

Causes as interventions

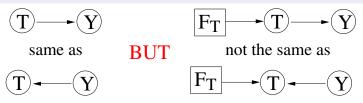
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w.r.t. conditional indep

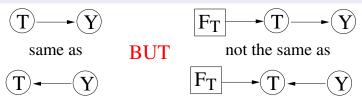
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■ LHS - T and Y are dependent

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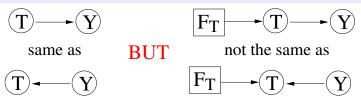
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- **RHS** top says $Y \perp F_T | T$ so T causes Y

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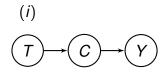
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- LHS T and Y are dependent
- **RHS** top says $Y \perp F_T | T$ so T causes Y
- **RHS** bottom says $Y \perp \perp F_T$ so T does not cause Y

DAGs for direct and indirect effects

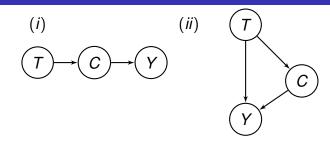


The DAGs encode conditional independence

(i) says $T \perp \!\!\!\perp Y | C$ i.e. if we know C, T cannot tell us anything more about Y

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DAGs for direct and indirect effects

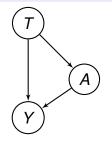


The DAGs encode conditional independence

- (i) says $T \perp | Y | C$ i.e. if we know C, T cannot tell us anything more about Y
- (ii) says nothing ... but that silence speaks volumes! We can extract more out of it if we add more variables see how later

Story

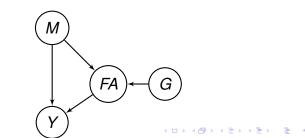
- A drug has strong head-aches as a side-effect
- Patients take aspirin to alleviate this
- It is thought that the aspirin might affect the outcome
- Drug company wants to know what are direct/indirect effects?



Example 2: meat and colon cancer

Story

- Meat consumption thought to be associated to colon cancer
- Also associated to colon cancer are genes that are responsible for breaking down fatty acids (CD36) via complex chemistry involving proteins etc
- Can we change/emulate gene behaviour
- for those with genes that predispose them to cancer?



Causal?

Remember Causes as interventions

- i.e. sex, age, race are not causes we cannot change them!
- treatments and other things we can in principle intervene on (education, weight) are potential causes

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 genes can be causes if we can conceive of intervening on them or their function

Causal?

Remember Causes as interventions

- i.e. sex, age, race are not causes we cannot change them!
- treatments and other things we can in principle intervene on (education, weight) are potential causes
- genes can be causes if we can conceive of intervening on them or their function

C can be intervened upon

- For there to be a direct/indirect effect, it must be possible to intervene on C
- Ex 1: aspirin intake of patients can in principle be controlled

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Ex 2: gene function changed or emulated

Identifiable?

Problem

- OK C can be intervened upon in principle....
- but in practice often not possible:
- cannot deny a patient with head-ache pain relief

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cannot intervene on a human (yet)

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Solution

- It all depends on the conditional independences!
- Using these and some cunning, we can identify the direct and indirect effects

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- cannot intervene on a human (yet)

Solution

- It all depends on the conditional independences!
- Using these and some cunning, we can identify the direct and indirect effects
- and often estimate them using data that are purely observational!

The type of direct and indirect effect you get depends on what you do to *C*

Ways of manipulating *C*

Fix C



The type of direct and indirect effect you get depends on what you do to *C*

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Ways of manipulating C

- Fix C
- Draw C from p(C|T)

The type of direct and indirect effect you get depends on what you do to *C*

Ways of manipulating C

Fix C

Draw C from p(C|T)

Draw C from an appropriate distribution

Each one of these manipulations can be described by a regime.

Ex 1: imaginary experiment I

Imagine you can fix aspirin intake e.g. everyone has to take no aspirins

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- Imagine you can fix aspirin intake e.g. everyone has to take no aspirins
- then can estimate the direct effect of treatment for no aspirins.
- Imagine you can randomise aspirins as well as treatment
- then can estimate the direct effect of treatment for 2 levels of aspirin and some kind of indirect effect by comparing the two direct effects.

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Problems

Not ethical!

Not realistic – people "out there" in the population do not behave like this.

Ex 1: imaginary experiment II

■ One random group of patients are given the control *F_T* = 0 and we record their "natural" aspirin intake *p*(*C*|*F_T* = 0)

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Ex 1: imaginary experiment II

- One random group of patients are given the control *F_T* = 0 and we record their "natural" aspirin intake *p*(*C*|*F_T* = 0)
- next we take the rest of the patients and
 - If a patient gets F_T = 1, they are administered aspirin according to p(C|F_T = 0)

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 If a patient gets F_T = 0, we leave them alone, and assume that they take aspirin according to p(C|F_T = 0)

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- If a patient gets F_T = 1, they are administered aspirin according to p(C|F_T = 0)
- If a patient gets F_T = 0, we leave them alone, and assume that they take aspirin according to p(C|F_T = 0)
- This way we can compare direct effects fixing not the aspirin intake, but its (control) distribution

Problem

Not ethical!

Ex 1: imaginary experiment III

- We cannot estimate p(C|T) lack of funds!
- So we pick a suitable distribution *p*(*D*) such that domain of *C* is same as *D*.
 - If a patient gets F_T = 1, they are administered aspirin according to p(D)

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If a patient gets F_T = 0, they are also administered aspirin according to p(D)

In this case, the comparison is not as easily interpretable, but we can still get a handle on the direct and indirect effects.

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- So we pick a suitable distribution *p*(*D*) such that domain of *C* is same as *D*.
 - If a patient gets F_T = 1, they are administered aspirin according to p(D)
 - If a patient gets F_T = 0, they are also administered aspirin according to p(D)
- In this case, the comparison is not as easily interpretable, but we can still get a handle on the direct and indirect effects.

Standardisation

The 2nd and 3rd regimes are forms of direct standardisation.

Formal definition of Regimes

Manipulating C

- Define *M_C* the variable of manipulations of *C*
- *M_C* has 4 regimes:
 - 1 $\mathbf{M}_{\mathbf{C}} = \emptyset$: *C* is left alone and comes from $p(C|T = t, F_T = \emptyset, M_C = \emptyset) = P_t$

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Formal definition of Regimes

Manipulating C

- Define M_C the variable of manipulations of C
- *M_C* has 4 regimes:
 - 1 $\mathbf{M_C} = \emptyset$: *C* is left alone and comes from $p(C|T = t, F_T = \emptyset, M_C = \emptyset) = P_t$
 - 2 $\mathbf{M}_{\mathbf{C}} = \mathbf{c}$: *C* is set to *c* with no uncertainty, distribution of *C* is δ_c

Formal definition of Regimes

Manipulating C

- Define M_C the variable of manipulations of C
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 - 3 $\mathbf{M}_{\mathbf{C}} = \mathbf{t}^*$: *C* is drawn from $p(C|F_T = \emptyset, M_C = t^*) = P_{t^*}$

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 - 2 $\mathbf{M}_{\mathbf{C}} = \mathbf{c}$: *C* is set to *c* with no uncertainty, distribution of *C* is δ_c
 - 3 $\mathbf{M}_{\mathbf{C}} = \mathbf{t}^*$: *C* is drawn from $p(C|F_T = \emptyset, M_C = t^*) = P_{t^*}$
 - 4 $M_C = D$: *C* is drawn from $p(D) = P_D$
- Note that Regime 2 is a special case of regime 3 which in turn is a special case of regime 4. Regime 1 is also a special case of 4.

Formal definition of Regimes continued

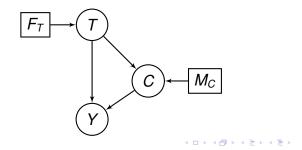
Conditional independences

()

$$\begin{array}{ccccc}
(Y,C) & \perp & F_T | T; \\
Y & \perp & M_C | (C,T); \\
C & \perp & (F_T,T) | M_C \neq \emptyset. \end{array} \tag{1}$$

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(3) only for regimes 2-4



The framework of these regimes is determined by *M_C* as defined

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- The framework of these regimes is determined by *M_C* as defined
- Adding other variables might require a redefinition of *M_C*
- As there are only 3 variables we consider here, we call this the 3DI framework
- Extensions are quite straightforward, might show some later

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Some simplyfing notation

$$E(Y|F_T = a, M_C = b) = E(a, b)$$

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$$E(Y|F_T = a, M_C = b) = E(a, b)$$

 $E(Y|F_T = a, M_C = b, C = c) = E(a, b, c)$
 $Pr(C = c|F_T = a, M_C = b) = p(c|a, b)$

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Definition of direct & indirect effects

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(4)
= $E(t, c) - E(t^*, c).$

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DE for *C* drawn from P_{t*} : Generated effect (*GDE*_{t*})

The direct effect of $F_T = t$ with respect to the baseline $F_T = t^*$ on response for *C* generated from P_{t^*} the distribution of *C* conditional on $F_T = t^*$ the baseline treatment is given by

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$$E(t, t^*) - E(t^*, t^*) \equiv E(t, t^*) - E(t^*, \emptyset).$$
(5)

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Definition of direct & indirect effects cont

DE for C drawn from P_D : Generated effect (GDE_D)

The direct effect of $F_T = t$ with respect to the baseline $F_T = t^*$ on response for *C* drawn from a specified distribution over *D* is given by:

$$E(t,D) - E(t^*,D) \tag{6}$$

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(6)

Total effect: TE

The total effect of $F_T = t$ with respect to the baseline $F_T = t^*$ on response is given by

$$E(Y|F_T = t) - E(Y|F_T = t^*).$$
 (7)

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 (7)

IDE for *C* drawn from $P (P \in \{\delta_c, P_{t^*}, P_D\})$ IDE = TE - DE

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Identifying effects in 3DI

- The TE is always identifiable, even when $F_T = \emptyset$ if there are no confounders
- If we can identify the DE then it follows we can identify the IDE, so focus is on DE
- We look first at scenario where there are no confounders and $F_T \neq \emptyset$
- Then consider what happens if there are confounders i.e. when

$$/ \perp M_C | (C, T)$$

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does not hold

• AND
$$F_T = \emptyset$$
.

$$GDE_{t^*} = \underbrace{E(t,t^*)}_{(a)} - \overbrace{E(t^*,\emptyset)}^{(b)}$$

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$$GDE_{t^*} = \underbrace{E(t, t^*)}_{(a)} - \underbrace{E(t^*, \emptyset)}_{(b)}$$
(b) is easy = $E(Y|T = t^*, F_T = \emptyset, M_C = \emptyset)$ by (1)
(a) =

$$=\sum_{c} E(t, t^*, c) \times p(c|t, t^*)$$
(8)

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$$=\sum_{c} E(t, \emptyset, c) \times p(c|\emptyset, t^*)$$
(8)
(9)

■ from (8)-(9) by (2) & (3)

$$GDE_{t^*} = \underbrace{E(t, t^*)}_{(a)} - \underbrace{E(t^*, \emptyset)}_{(E(t^*, \emptyset))}$$
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(8)
$$= \sum_{c} E(t, \emptyset, c) \times p(c|\emptyset, t^*)$$
(9)
$$= \sum_{c} E(Y|T = t, F_T = \emptyset, M_C = \emptyset, C = c)$$

$$\times p(C = c|T = t^*F_T = \emptyset, M_C = \emptyset)$$
 (10)

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from (8)-(9) by (2) & (3)
from (9)-(10) by (1)

Pros

So it is possible to identify and hence estimate the 3DI effects from data that are purely observational for both *T* and *C* as *F*_T and *M*_C are both Ø.

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Simple formulae, easy to understand

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- Simple formulae, easy to understand once you get the hang of it

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- Simple formulae, easy to understand once you get the hang of it

Cons

No counfounder assumption is unlikely to hold in practice

Confounder between *T* and *C* or *T* and *Y* This is only really an issue if $F_T = \emptyset$

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Confounder between T and C or T and Y

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Confounder between C and Y

This requires a redefiniton of M_C as the distribution of C now depends on this confounder (call it W) as well as T.

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W is a niusance

- Ex 1: Two doctors administer the treatment.
- One knows that the side-effect is headache, tells patients not to take aspirin
- The other knows nothing so his patients are ignorant

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- In this situation we want to get rid of W

W is relevant

Ex 1: W is sex: women suffer more side-effect than men.

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The two roles of W need to be treated in slightly different ways.

W as sex

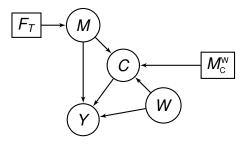
Just consider this - the other case in the paper New variable and conditional independence - 4DI

Define M_c^w as a manipulation variable on C s.t.

$$(Y, C, W) \perp F_T | T$$
(11)

$$W \perp (F_T, M_c^w)$$
(12)

$$Y \perp M_{c}^{W} | (T, C, W).$$
 (13)



W as sex continued

Regimes of *M*^w_c

 M^w_c = Ø: C is left alone
 M^w_c = t^{*}: C is drawn from p(C|T = t^{*}, W = w, F_T = Ø) where w corresponds to the realised value of W.

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Regimes of M_c^w

- 1 $M_{c}^{w} = \emptyset$: *C* is left alone
- 2 $M_c^w = t^*$: *C* is drawn from $p(C|T = t^*, W = w, F_T = \emptyset)$ where *w* corresponds to the realised value of *W*.
- No need for other regimes as they sever relationship between C and W and are equivalent to 3DI
- that means that there is only 1 extra DE
- Regime 2 induces

$$C \perp (F_T, T) | M_c^w = t^*$$
(14)

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$$E(t, t^*, W = 1) - E(t^*, t^*, W = 1)$$

$$E(t, t^*, W = 1) - E(t^*, t^*, W = 1)$$

= $\sum_{c} [E(t, \emptyset, c, W = 1)]$

$$E(t, t^*, W = 1) - E(t^*, t^*, W = 1)$$

= $\sum_{c} [E(t, \emptyset, c, W = 1) \times p(c|t^*, \emptyset, W = 1)]$

$$E(t, t^*, W = 1) - E(t^*, t^*, W = 1) = \sum_{c} [E(t, \emptyset, c, W = 1) \times p(c|t^*, \emptyset, W = 1)] - E(t^*, \emptyset, W = 1).$$

$$E(t, t^*, W = 1) - E(t^*, t^*, W = 1) = \sum_{c} [E(t, \emptyset, c, W = 1) \times p(c|t^*, \emptyset, W = 1)] - E(t^*, \emptyset, W = 1).$$

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This can be further expressed with $F_{T} = \emptyset$ and can thus also be estimated from purely observational data.

Conclusions

- DAGs and conditional independences are useful for expressing D&I effects
- We need to be clear about what is going on though what Cl's hold?
- Provided we do this, we have a wealth of DEs we can estimate
- Also remember that the concepts don't make much sense if we cannot intervene on the intermediate at all.
- The 3DI and 4DI are tools that can be used to clarify what we mean.
- The framework can be extended further.

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