# On the identification of discrete graphical models with hidden nodes 

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## Introduction

- The focus is on discrete graphical models with one latent binary variable
- Conditions of identification of these models are established
- They are based on the structure of the conditional independence graph


## The model I

Let $G^{K}=(K, E)$ be an undirected graph with:

- node set $K=\{0,1, \ldots, n\}$
- edge set $E=\{(i, j)\}$, whenever vertices $i$ and $j$ are adjacent in $G^{K}$

Partition $K=\{0, O\}$, with $O=\{1, \ldots, n\}$.

## The model II

To each node $v$ is associated a discrete random variable $A_{v}$. A discrete undirected graphical model is:

- a family of joint distributions of the variables $A_{v}, v \in K$, satisfying the Markov property with respect to $G^{K}$,
- namely that $A_{v} \Perp A_{u} \mid$ rest whenever $u$ and $v$ are not adjacent in $G^{K}$.

Let $A_{0}$ be a binary unobserved variable. Let $A_{v}$ with $v \in O$ be the observable random variables.

We consider a multidimensional contingency table obtained by the cross classification of $N$ objects according to $A_{v}, v \in K$.


Figure: An example of c.i. graph

Let $X$ be the $2 \ell \times 1, \ell=\prod_{v=1}^{n} I_{v}$, vector of the counts of the contingency table, with $A_{0}$ running the slowest.

## The model III

## We assume

- that the elements of $X$ are independent Poisson random variables with $E(X)=\mu_{X}$.
- $\log \mu_{X}=Z \beta$, where $Z$ is a $2 \ell \times p$ design matrix such that the joint distribution of $A_{v}, v \in K$, factorizes according to $G^{K}$
- $\beta$ is a p-dimensional vector of unknown parameters. We adopt the corner point parametrization that takes as first level the cell with $A_{v}=0$, for all $v \in K$.

The multinomial case can be addressed by assuming $N$ fixed.

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## The observable r.v.'s

Let $Y$ the $\ell \times 1$ vector of the counts in the marginal table, obtained by the cross classification of the $N$ objects according to the observed variables only.

- $Y$ may be written as $Y=L X$, with $L=(1,1) \otimes I_{\ell}$.
- each element of $Y$ is Poisson, with $E(Y)=\mu_{Y}=L e^{Z \beta}$

For now, we assume that also the observable variables are binary (can be removed).

Note that the typical element of $\mu_{Y}$ is

$$
e^{a}+e^{a+b}
$$

For instance (with reference to the previous example):

$$
\mu\left[\begin{array}{llll}
1 & 0 & 0 & 0
\end{array}\right]=\underbrace{e^{\mu_{0}+\beta_{1}}}_{\mu\left[\begin{array}{lll}
0 & 1 & 0
\end{array}\right]}+\underbrace{e^{\mu_{0}+\beta_{1}+\beta_{0}+\beta_{01}}}_{\mu\left[\begin{array}{lll}
1 & 1 & 0
\end{array} 00\right]}
$$

## An example

We will assume that all observable r.v.'s are connected to the latent one (if not take the subgraph of the relevant var's).


Figure: Example of (a) a $G^{K}$ graph (b) the induced subgraph $G^{O}$ and (c) $\bar{G}^{O}$

## Local identification - I

By the inverse function theorem, a model is locally identified if the rank of the transformation from the natural parameters $\mu_{Y}$ to the new parameters $\beta$ is full. This is equivalent to the rank of following derivative matrix

$$
\begin{equation*}
D(\beta)^{T}=\frac{\partial \mu_{Y}^{T}}{\partial \beta}=\frac{\partial\left(L e^{Z \beta}\right)^{T}}{\partial \beta}=(L R Z)^{T} \tag{1}
\end{equation*}
$$

being full, where $R=\operatorname{diag}\left(\mu_{X}\right)$. Note that the $(i, j)$-th element of $D(\beta)$ is the partial derivative of the $i$-th component of $\mu_{Y}$ with respect to $\beta_{j}$ the $j$-th element of $\beta$.

## Local identification - II

Note that, by setting $t_{j}=e^{\beta_{j}}$ for any parameter $\beta_{j}$, the parametrization map turns into a polynomial one.

This implies that if there exists a point in the parameter space of $t_{j}$ at which the Jacobian has full rank, then the rank is full almost everywhere.

Therefore, either (a) there is no point in the parameter space where the rank is full or (b) the rank is full almost everywhere.

## Object of the paper:

(a) to establish a necessary and sufficient condition on the graphical model to be full-rank almost everywhere in the parameter space
(b) to give the expression of the subspace where identifiability breaks down.

## The general idea

Conditions are established such that (a) there is an ordering of rows and columns of $D(\beta)$, such that:

$$
\mathrm{D}(\beta)=
$$


and (b) blocks $A_{i}$ are full-rank.

## Condition (a): the existence of the ordering

It requires that the set $O$ of the observed variables contains at least three variables.

Moreover, the complementary graph of the sub-graph $G^{O}$ is connected, with an $m$-clique, $m \geq 3$.

## Condition (b): the blocks $A_{i}$ are full rank

Consider $I \subseteq O$, let $\mu_{I}$ be element of $\mu_{Y}$ associated to the entry of the contingency table having value zero for all variables except the variables in $I$.

Let $\beta_{l}$ be the term of interactions among the variables in $I$.
Let $d_{l}$ be the row of the matrix $D(\beta)$ corresponding to the first order partial derivative of $\mu_{l}$ with respect of $\beta$.

## The generic $A_{i}$

Let $S$ be a complete subgraph of $G^{O}$ and $S^{\prime} \supset S$ :

|  | $\beta_{S}$ | $\beta_{\{0, S\}}$ |
| :---: | :---: | :---: |
| $\mathrm{d}_{S}$ | $e^{a}\left(1+e^{b}\right)$ | $e^{a+b}$ |
| $\mathrm{~d}_{S^{\prime}}$ | $e^{a+a^{\prime}}\left(1+e^{b+b^{\prime}}\right)$ | $e^{a+a^{\prime}+b+b^{\prime}}$ |

$A_{i}$ is not full rank if and only if $b^{\prime}=0$. Therefore, conditions to assure that $b^{\prime} \neq 0$ are given (see the paper).

What are the possible expressions of $b^{\prime}$ ? Here are some example:

1. $b^{\prime}=\beta_{\{0, v\}}$ with $v \in O$ a single node.
2. $b^{\prime}=\beta_{\{0, I\}}$ with $\{0, I\}$ complete in $G^{K}$.
3. $b^{\prime}$ is a sum of different elements of $\beta$ (e.g $\left.\beta_{0}+\beta_{01}\right)$.

If we assume that the graph is faithfull, then 1. and 2. lead to a full rank sub-matrix $A_{i}$. Conditions are given to avoid 3 ., based on the c.i. graph (see the paper).

## Violations of the conditions (a) and (b)

Violation of condition (b) leads to the expression of the subspace of null measure where identifiability breaks down.

Violation of condition (a) leads to a non full-rank model everywhere in the parameter space.

Here inspection of the matrix $D(\beta)$ is necessary to reveal (a) the rank of $D(\beta)$ and (b) which columns of $D(\beta)$ are linearly dependent.

## An example of identified model



Figure: Full rank model, rank of $D(\beta)=13$

The LRT against the saturated model has $\chi^{2}(2)$ as asymptotic distribution.

## An example of not identified model



Figure: Violates condition (a), rank of $D(\beta)=11$

The true model dimension is 11 .

## Focus on the model

Inspection of $D(\beta)$ leads to the following conclusions:
columns $\beta_{12}, \beta_{012}, \beta_{34}, \beta_{034}$ are aliased (and so are other four parameters);
if we impose $\beta_{012}=\beta_{034}=0$ the model has full rank.
The LRT against the saturated model seems to have a $\chi^{2}(4)$ as asymptotic distribution (instead of a $\chi^{2}(2)$ ).

## Histograms of p-value

Histograms 2500 p-values of LRT

$p$-values taken from chi-square $\mathrm{df}=4$

Histograms 2500 p-values of LRT


## Conclusions

- We have derived conditions for the proposed model to be locally identified almost everywhere.
- Expression of the null measure non identified subspace has been derived.
- For non full rank models, a way to compute the true model dimension has been established.
- This leads to compute correctly the asymptotic distribution of the LRT against the saturated model (work in progress, with A. Forcina).


## Research question

## Will algebraic statistics make may life easier?

