

Continuous-time Importance Sampling for Multivariate  
Diffusion Processes  
(Avoiding time-discretisation approximation error)

Paul Fearnhead

Lancaster University

`www.maths.lancs.ac.uk/~fearnhea`

Part of ongoing work with Krys Latuszynski, Gareth Roberts and  
Giorgos Sermaidis

## Diffusions

A **diffusion** is a **continuous-time Markov process** with **continuous** sample paths. We can define a diffusion as the solution of a **Stochastic Differential Equation (SDE)**:

$$dX_t = \mu(X_t)dt + \sigma(X_t)dB_t.$$

Intuitively this defines the **dynamics** over small time intervals. **Approximately** for small  $h$ :

$$X_{t+h}|X_t = x_t \sim x_t + h\mu(x_t) + h^{1/2}\sigma(x_t)Z,$$

where  $Z$  is a standard normal random variable.

## Transition Densities

We will denote the [transition density](#) of the diffusion by

$$p(y|x, h) = p(X_{t+h} = y | X_t = x).$$

It satisfies Kolmogorov's forward equation:

$$\frac{\partial}{\partial t} p(y|x, t) = \mathcal{K}_y p(y|x, t),$$

for some [forward-operator](#)  $\mathcal{K}_y$  which acts on  $y$ .

Generally the transition density is [intractable](#). Exceptions include models with both  $\sigma(x) = \sigma$  and  $\mu(x) = a + bx$ .  $X_t$  is then a [Gaussian process](#).

## The Exact Algorithm

Generally simulation and inference for diffusions is performed by [approximating](#) the diffusions by a [discrete-time Markov process](#).

However, work by [Beskos, Papaspiliopoulos and Roberts](#) demonstrate how to simulate from a [class](#) of diffusion models where (possibly after transformation):

- The volatility is the identity:  $\sigma(x) = \mathbf{I}$ .
- The drift is the gradient of a potential:  $\mu(x) = \nabla A(x)$ .

This can be applied to [almost all](#)  $1-d$  diffusions, and [almost no](#) others.

## Current Approaches: The Exact Algorithm

The exact Algorithm is a [Rejection Sampler](#) based on proposing paths from Brownian motion.

The acceptance probability for the path is (for  $\sigma(x) = 1$ ) proportional to:

$$\begin{aligned} & \exp \left\{ - \int_0^T \mu(X_t) dX_t + \frac{1}{2} \int_0^T \mu(X_t)^2 dt \right\} \\ & = \exp \left\{ A(X_T) - A(X_0) - \frac{1}{2} \int_0^T (\mu(X_t)^2 + \mu'(X_t)) dt \right\}. \end{aligned}$$

Whilst this cannot be evaluated, events with this probability can be simulated.

## Avoiding time-discretisation Errors: Why?

Beskos, Papaspiliopoulos, Roberts and Fearnhead (2006) extend the rejection sampler to an importance sampler, and show how this can be used to perform inference for diffusions which avoids time-discretisation approximations.

Why may these methods be useful?

- Error in estimates are purely Monte Carlo. Thus it is easier to quantify the error.
- Time-discretisation may tend to use substantially finer discretisations than are necessary: possible computational gains?
- Error is  $O(C^{-1/2})$ , where  $C$  is CPU cost. Alternative approaches have errors that are e.g.  $O(C^{-1/3})$  or worse.

## Our Aim

Our aim was to try and extend the ability to perform inference without time-discretisation approximations to a wider class of diffusions.

The key is to be able to unbiasedly estimate expectations, such as  $\mathbf{E}(f(X_t))$  or  $\mathbf{E}(f(X_{t_1}, \dots, X_{t_m}))$ .

## The Exact Algorithm: Generalising Conditions

The condition  $\mu(x) = \nabla A(x)$  is required to replace the stochastic integral by a Lebesgue one. It is a **necessary and sufficient** condition for Girsanov's formula to be bounded for bounded sample paths.

The condition  $\sigma(x)$  is the identity as otherwise we do not have a proposal distribution that is **tractable** and **absolutely continuous** wrt to the target:

Consider two diffusions with **different** diffusion coefficients,  $\sigma_1$  and  $\sigma_2$ , then their laws as **NOT** mutually absolutely continuous ...

even though their finite-dimensional distributions typically are.



## New Approach: CIS

We now derive a continuous-time importance sampling (CIS) procedure for unbiased inference for general continuous-time Markov models.

We will describe the CIS algorithm for generating a single realisation. So at any time  $t$  we will have  $x_t$  and  $w_t$ , realisations of random variables  $X_t, W_t$  such that

$$\mathbb{E}_p(f(X_t)) = \mathbb{E}_q(f(X_t)W_t).$$

The former expectation is wrt to the target diffusion, the latter wrt to CIS procedure.

We will use a proposal process with tractable transition density  $q(x|y, t)$  (and forward-operator  $\mathcal{K}_x^{(1)}$ ).

## A discrete-time SIS procedure

First consider a discrete-time SIS method aimed at inference at times  $h, 2h, 3h, \dots$ .

- (0) Fix  $x_0$ ; set  $w_0 = 1$ , and  $i = 1$ .
- (1) Simulate  $X_{ih} = x_{ih}$  from  $q(x_{ih}|x_{(i-1)h})$ .
- (2) Set

$$w_i = w_{i-1} \frac{p(x_{ih}|x_{(i-1)h}, h)}{q(x_{ih}|x_{(i-1)h}, h)}$$

- (3) Let  $i = i + 1$  and goto (1).

**Problems:** cannot calculate weights, and often the efficiency degenerates as  $h \rightarrow 0$  for fixed  $T$ .

## Random weight SIS

It is valid to replace the weight in the SIS procedure by a **random variable** whose expectation is equal to the weight.

A simple way to do this here is to define

$$r(y, x, h) = 1 + \left( \frac{p(y|x, h)}{q(y|x, h)} - 1 \right) \frac{1}{\lambda h},$$

and introduce a **Bernoulli** random variable  $U_i$ , with success probability  $\lambda h$ .

Then

$$\frac{p(y|x, h)}{q(y|x, h)} = \mathbb{E} \{ (1 - U_i) \cdot 1 + U_i r(y, x, h) \}.$$

## Random weight SIS

Now we can have a [random weight SIS](#) algorithm:

- (0) Fix  $x_0$ ; set  $w_0 = 1$ , and  $i = 1$ .
- (1) Simulate  $X_{ih} = x_{ih}$  from  $q(x_{ih}|x_{(i-1)h})$ .
- (2) Simulate  $U_i$ . If  $U_i = 1$  then set  $w_i = w_{i-1}r(x_{ih}, x_{(i-1)h}, h)$ , otherwise  $w_i = w_{i-1}$ .
- (3) Let  $i = i + 1$  and [goto \(1\)](#).

This is a less efficient algorithm than the previous one, but it enables us to now use two tricks: [retrospective sampling](#) and [Rao-Blackwellisation](#).

## Retrospective Sampling

We only need to update the weights at time-points where  $U_i = 1$ . At these points we need to simulate  $X_{ih}, X_{(i-1)h}$  to calculate the new weights.

If  $j$  is the most recent time when  $U_j = 1$ , then the distribution of  $X_{ih}$  is given by  $q(x_{ih}|x_{jh}, (i - j)h)$ .

Given  $x_{jh}$  and  $x_{ih}$  the conditional distribution of  $X_{(i-1)h}$  is

$$q(x_{(i-1)h}|x_{jh}, x_{ih}) = \frac{q(x_{(i-1)h}|x_{jh}, (i - j - 1)h)q(x_{ih}|x_{(i-1)h}, h)}{q(x_{ih}|x_{jh}, (i - j)h)}.$$

## New SIS algorithm

Using these ideas we get:

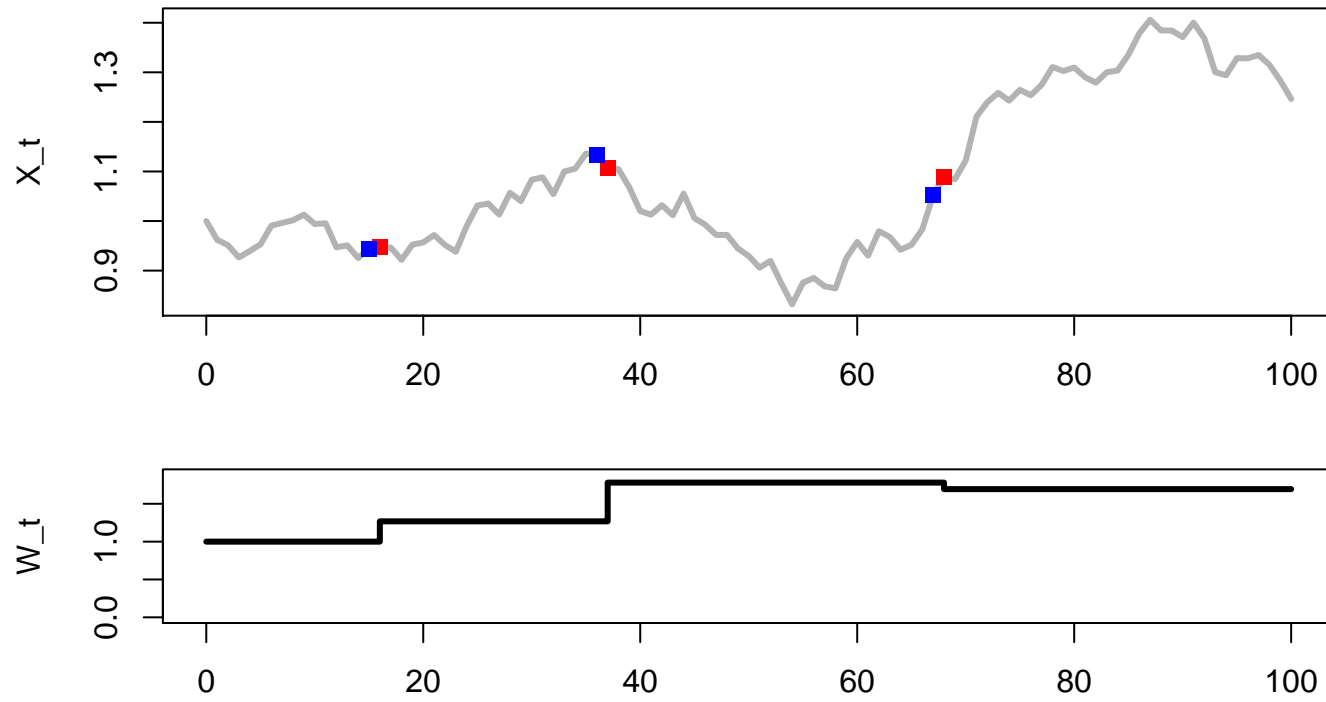
- (0) Fix  $x_0$ ; set  $w_0 = 1$ ,  $j = 0$  and  $i = 1$ .
- (1) Simulate  $U_i$ ; if  $U_i = 0$  goto (3).
- (2) [ $U_i = 1$ ] Simulate  $X_{ih}$  from  $q(x_{ih}|x_{jh}, (i-j)h)$  and  $X_{(i-1)h}$  from  $q(x_{(i-1)h}|x_{jh}, x_{ih})$ .  
Set

$$w_i = w_j r(x_{ih}, x_{(i-1)h}, h).$$

- (3) Let  $i = i + 1$  and goto (1).

If we stop the SIS at a time point  $t$ , then  $X_t$  can be drawn from  $q(x_t|x_{jh}, t - jh)$ ; and the weight is  $w_j$ .

# Example



## Rao-Blackwellisation

At time  $ih$ , the incremental weight depends on  $x_{ih}$  and  $x_{(i-1)h}$ . Rather than simulating both we simulate  $x_{ih}$ , and use an expected incremental weight

$$\rho_h(x_{ih}, x_{jh}, (j-i)h) = \mathbb{E} \left( r(x_{ih}, X_{(i-1)h}, h) \mid x_{jh} \right),$$

with expectation with respect to the conditional distribution of  $X_{(i-1)h}$  given  $x_{jh}, x_{ih}$  under the proposal:

$$\mathbb{E} \left( r(x_{ih}, X_{(i-1)h}, h) \mid x_{jh} \right) = \int r(x_{ih}, x_{(i-1)h}, h) q(x_{(i-1)h} \mid x_{jh}, x_{ih}) dx_{(i-1)h}.$$



## New SIS algorithm

Using these ideas we get:

- (0) Fix  $x_0$ ; set  $w_0 = 1$ ,  $j = 0$  and  $i = 1$ .
- (1) Simulate  $U_i$ ; if  $U_i = 0$  goto (3).
- (2) [ $U_i = 1$ ] Simulate  $X_{ih}$  from  $q(x_{ih}|x_{jh}, (i - j)h)$  and set

$$w_i = w_j \rho_h(x_{ih}, x_{jh}, (i - j)h).$$

- (3) Let  $i = i + 1$  and goto (1).

If we stop the SIS at a time point  $t$ , then  $X_t$  can be drawn from  $q(x_t|x_{jh}, t - jh)$ ; and the weight is  $w_j$ .

## Continuous-time SIS

The previous algorithm cannot be implemented as we do not know  $p(\cdot|\cdot, h)$ . However, if we consider  $h \rightarrow 0$  we obtain a **continuous-time** algorithm that can be implemented.

The **Bernoulli process** converges to a **Poisson-process**.

In the limit as  $h \rightarrow 0$ , if we fix  $t = ih$  and  $s = jh$  we get

$$\rho(x_t, x_s, t - s) = \lim_{h \rightarrow 0} \rho_h(x_t, x_s, t - s) = 1 + \frac{1}{\lambda} \left( \frac{(\mathcal{K}_x - \mathcal{K}_x^{(1)})q(x|x_s, t - s)}{q(x|x_s, t - s)} \right) \Big|_{x=x_t} .$$

## CIS Algorithm

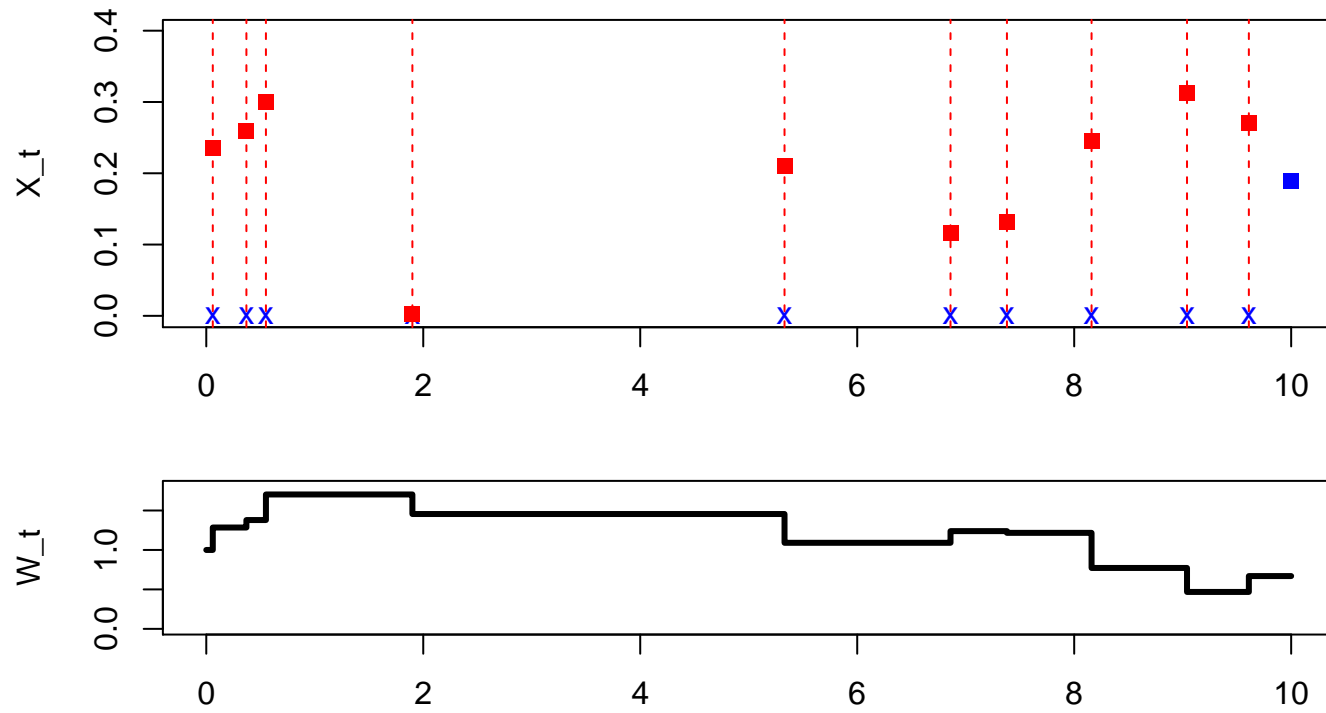
- (0) Fix  $x_0$ ; set  $w_0 = 1$  and  $s = 0$ .
- (1) Simulate the time  $t$  of the next event after  $s$  in a Poisson process of rate  $\lambda$ .
- (2) Simulate  $X_t$  from  $q(x_t|x_s, t - s)$ ; and set

$$w_t = w_s \times \rho(x_t, x_s, t - s).$$

- (3) Goto (1).

If we stop the SIS at a time point  $T$ , then  $X_T$  can be drawn from  $q(x_T|x_s, T - s)$ ; and the weight is  $w_j$ .

# Example CIS



## Does it work?

Not always! A necessary and sufficient condition for the method to be valid (ie unbiased) is that the weight process  $\{w_s; s \geq 0\}$  is a martingale.

This does not automatically happen as  $W_t$  can be negative: thus  $E(|W_t|)$  can be infinite.

## CIS: Implementation for diffusions

The **target** process is

$$dX_t = \mu(X_t)dt + \sigma(X_t)dB_t.$$

Denote event times by  $\tau_1, \tau_2, \dots$ , and  $\tau(t)$  the time of the most-recent event prior to  $t$ .

Proposal process needs to be tractable: e.g. constant drift and volatility.

- Can allow rate of Poisson process to depend on time since last event:  $\lambda = \lambda(t - \tau(t))$ .
- At each renewal, can **update** the importance process:

$$dX_t = b(\tau_i)dt + v(\tau_i)dB_t.$$

## Does it work?

In **almost all** cases where the proposal is **not** chosen to have  $v(\tau_i) = \sigma(X_{\tau_i})$  then the weight process turns out to **NOT** be in  $L^1$ !

What about the **copycat scheme**?  $v(\tau_i) = \sigma(X_{\tau_i})$ ,  $b(\tau_i) = \mu(X_{\tau_i})$

### **Theorem:**

1. If  $\sigma$  and  $\mu$  are globally Lipschitz, and  $\sigma$  is bounded away from 0, then the copycat scheme is **valid**.
2. For all  $p > 1$ , there exists  $\epsilon > 0$  such that choosing  $\lambda(u) \propto u^{-1+\epsilon}$  ensures that  $\{w_s, s \geq 0\}$  is an  $L^p$  martingale.

## Example: CIR Diffusion

We consider estimating the transition density for a 2-d CIR model:

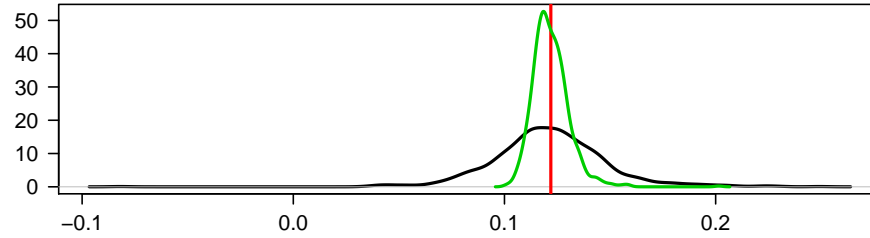
$$\begin{bmatrix} dX_t^{(1)} \\ dX_t^{(2)} \end{bmatrix} = \begin{bmatrix} -\rho_1(X_t^{(1)} - \mu_1) \\ -\rho_2(X_t^{(2)} - \mu_2) \end{bmatrix} dt + \begin{bmatrix} \sigma_1 \sqrt{X_t^{(1)}} & 0 \\ \rho\sigma_2 \sqrt{X_t^{(2)}} & \sigma_2 \sqrt{(1 - \rho^2)X_t^{(2)}} \end{bmatrix} \begin{bmatrix} dB_t^{(1)} \\ dB_t^{(2)} \end{bmatrix}$$

We compare the CIS with a time-discretisation approach based on the ideas in Durham and Gallant (2002), for varying CPU cost.

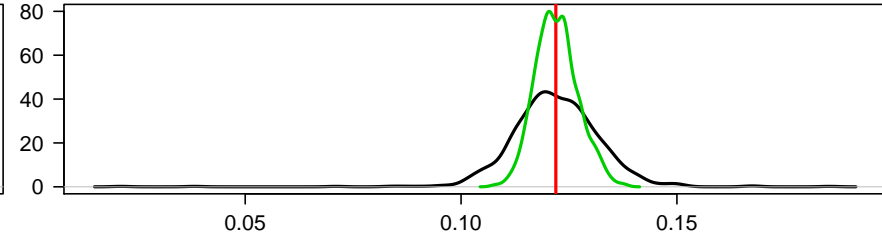


# Example: CIR Diffusion

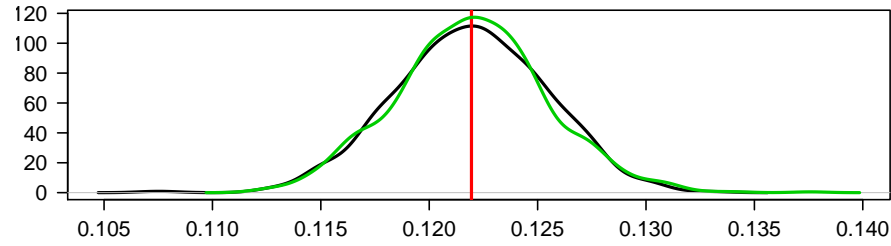
Computational setting 1



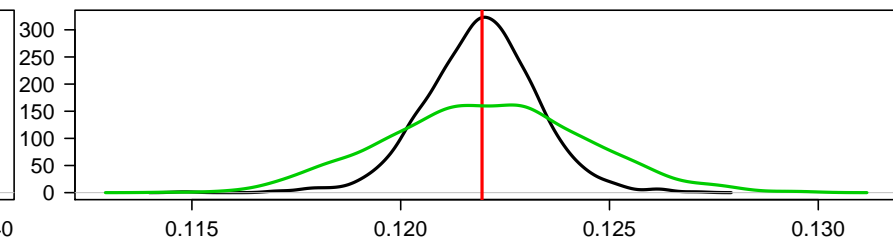
Computational setting 2



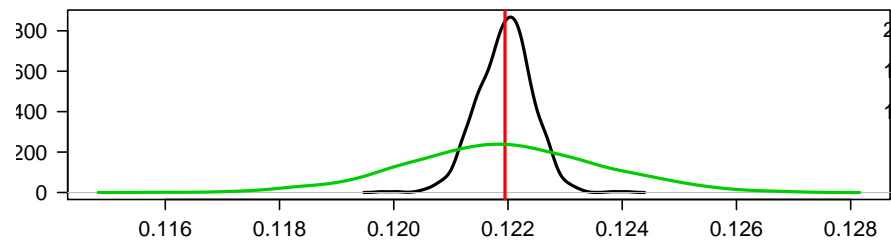
Computational setting 3



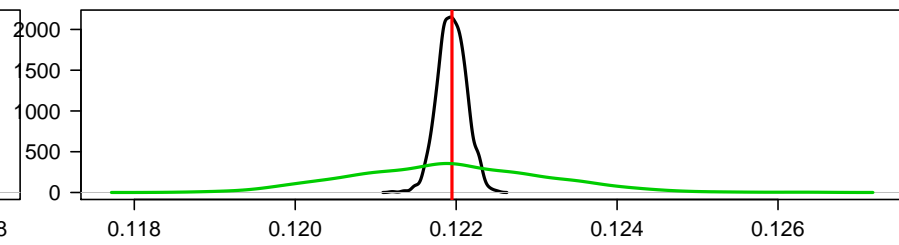
Computational setting 4



Computational setting 5



Computational setting 6



## Example: Hybrid Systems

CIS can be applied to other [continuous-time Markov](#) processes.

One example is a [hybrid](#) linear diffusion/Markov-jump process:

$$dX_t = (a(t, Y_t) + b(t, Y_t)X_t) dt + \sigma(t, Y_t)dB_t,$$

and  $Y_t$  is a Markov-jump process with generator (rate-matrix)  $Q(X_t)$ .

Such processes arise in [systems biology](#) and [epidemic models](#)

## Example: Hybrid Systems

If we can bound the rate,  $\lambda(X_t, y_t)$  of leaving a state  $y_t$  by  $\bar{\lambda}$ , then we can simulate from this process using **thinning**:

- Simulate the next time,  $\tau$  from a **Poisson Process** with rate  $\bar{\lambda}$ .
- Simulate  $X_\tau$ .
- With probability  $\lambda(X_\tau, y_t)/\bar{\lambda}$  **simulate an event** in the  $Y_t$  process.

**CIS** can be implemented in a way similar to thinning, but **does not require a bound**,  $\bar{\lambda}$ . Instead if  $\lambda(X_\tau, y_t) > \bar{\lambda}$  we get an **Importance Sampling Correction**.

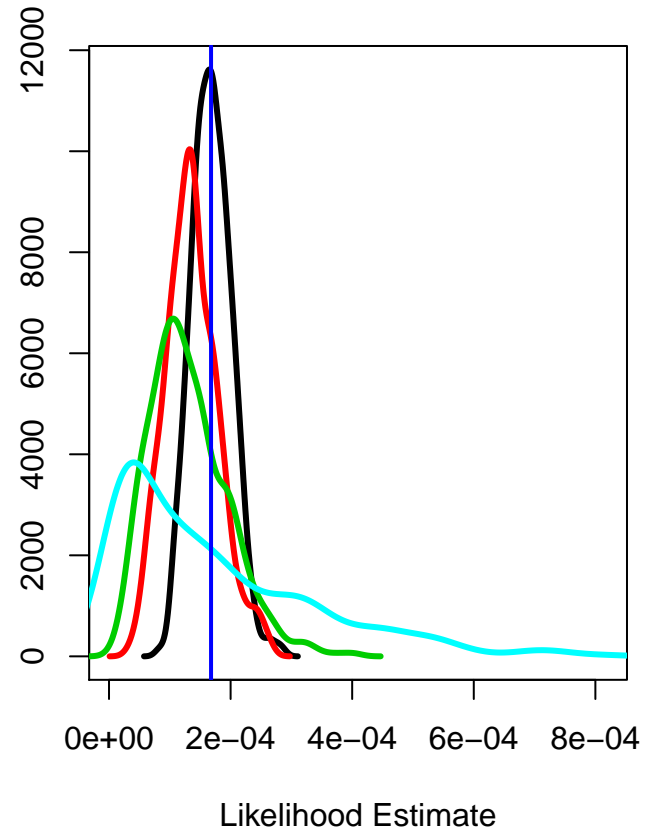
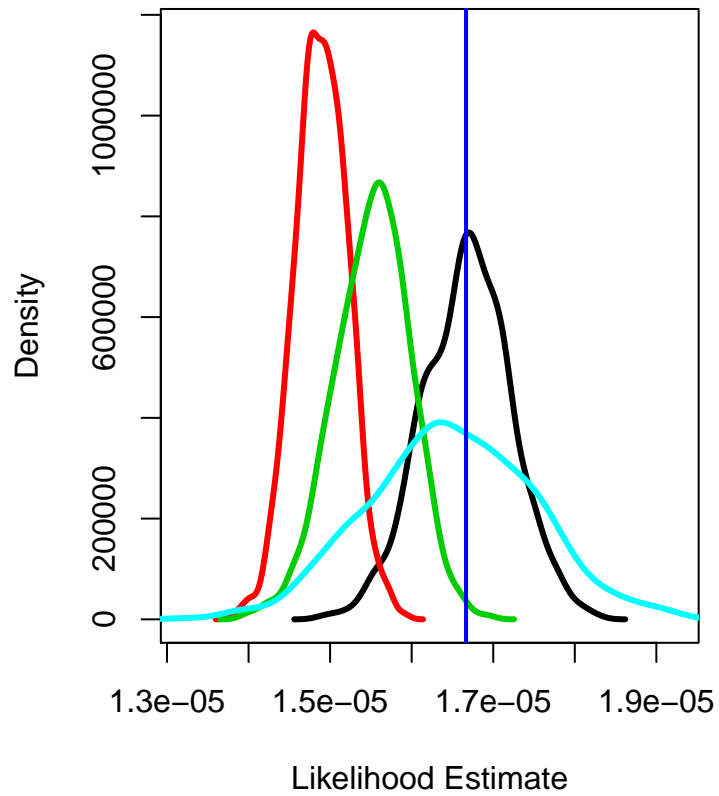
## Auto-Regulatory System

We applied this to a hybrid system based on a 4-dimensional model of an [autoregulatory system](#).

We looked at the accuracy of estimating the likelihood of data at a single time-point.

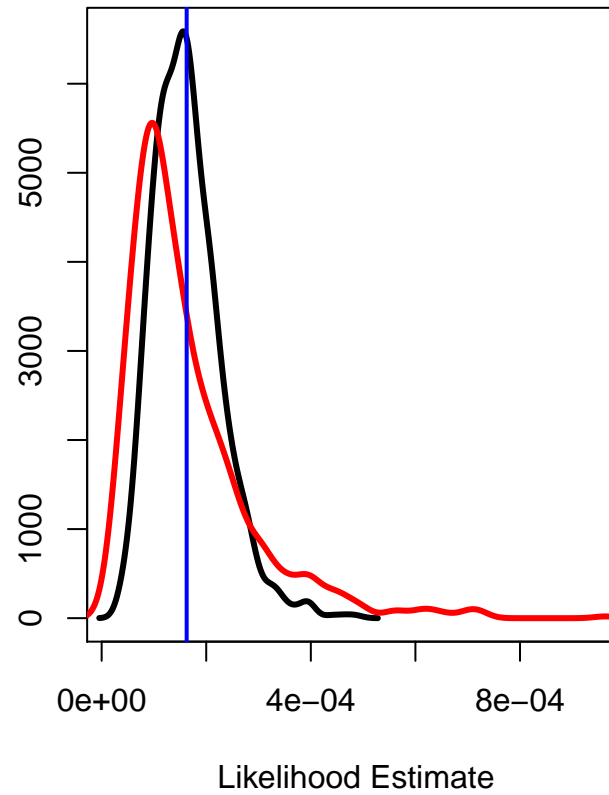
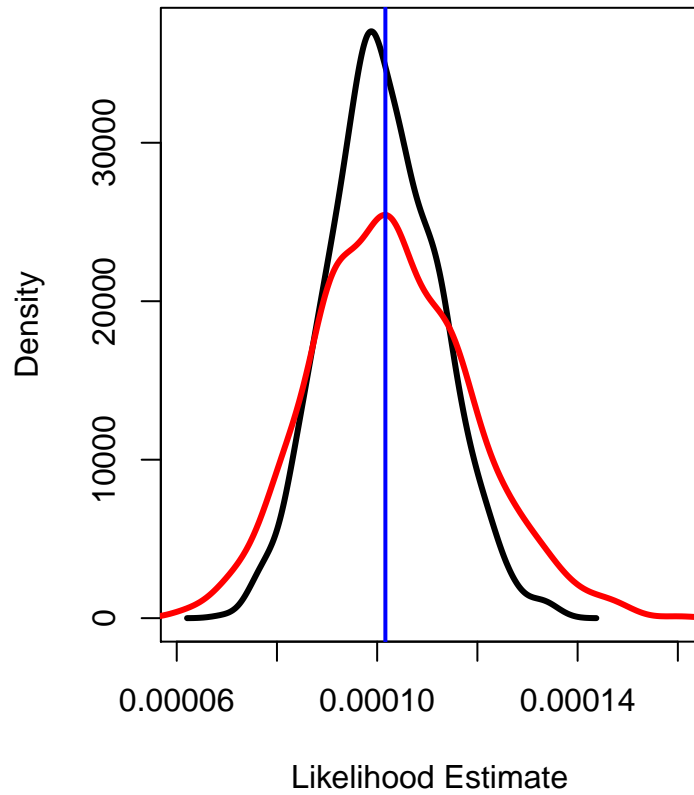
We utilised the tractability of the  $X_t$  process after the last event-time at which we (potentially) updated the  $Y_t$  process to improve the accuracy of our estimate – this advantages methods with fewer event times.

# Auto-Regulatory System: Comparison with Euler



## Comparison with (approximate) Thinning

Thinning with bound on rates chosen so that  $\Pr(\lambda(X_\tau, y_t) < \bar{\lambda}) \approx 1$



## Discussion

This is a very [flexible](#) and [potentially](#) powerful method. Can be used to unbiasedly estimate density (likelihood), expectations, etc.

There are [numerous variance reduction methods](#) that can be used

There is a related approach for diffusions by [Wolfgang Wagner](#). His approach can be viewed as [Importance Sampling](#), whereas ours is most similar to [Sequential Importance Sampling](#). This has advantages in terms of using ideas (resampling, adapting proposals) from SIS to improve accuracy.

There are links of our method with [Thinning](#) of Jump-Markov processes.

Dealing with the negative weights is an important issue.