# Terrorists never congregate in even numbers<sup>1</sup>

(or: Some strange results in fragmentation-coalescence )

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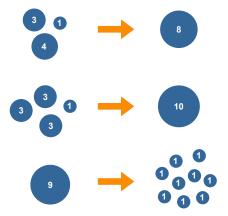
<sup>&</sup>lt;sup>1</sup>Joint work with Steven Pagett, Tim Rogers



- Consider a collection of n identical particles (terrorists/opinions), grouped together into some number of clusters (cells/consensus). We define a stochastic dynamical process as follows:
- Every k-tuple of clusters coalesces at rate  $\alpha(k)n^{1-k}$ , independently of everything else that happens in the system. The coalescing cells are merged to form a single cluster with size equal to the sum of the sizes of the merged clusters.
- Clusters fragment (terrorist cells are dispersed/consensus breaks) at constant rate  $\lambda > 0$ , independently of everything else that happens in the system. Fragmentation of a cluster of size  $\ell$  results in  $\ell$  'singleton' clusters of size one.

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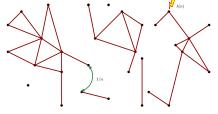


- Without fragmentation, the model falls within the domain of study of Smoluchowski coagulation equations, originally devised to consider chemical processes occurring in polymerisation, coalescence of aerosols, emulsication, flocculation.
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## Model history (but only for dyadic coalescence)

- This model is a variant of the one presented in: Bohorquez, Gourley, Dixon, Spagat & Johnson (2009) Common ecology quantifies human insurgency Nature 462, 911-914.
- It is also related to: Ráth and Tóth (2009) Erdős-Rènyi random graphs + forest fires = self-organized criticality, 14 Paper no. 45, 1290-1327.



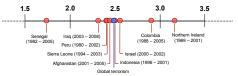
In a system of size n 'vacant' edges become 'occupied' at rate 1/n, each site 'hit by lightning' at rate  $\lambda(n)$  annihilating to singletons the cluster in which it is contained.

### Heavy-tailed terrorism

- In the insurgency model, two blocks merge if a terrorist in each block make a connection, which they do at a fixed rate. This means that coalescence is more likely for a big terrorist cell.
- The macroscopic-scale, large time limit of the insurgency model for a "slow rate of fragmentation" shows that the distribution of block size is heavy tailed:

"P(typical block = 
$$x$$
)  $\approx$  const.  $\times x^{-\alpha}$ ,  $x \to \infty$ ."

 Taken from Bohorquez, Gourley, Dixon, Spagat & Johnson (2009):



## Back to our model: Generating function

- For each  $n \in \mathbb{N}$ , and  $k \in \{1, ..., n\}$ , the state of the system is specified by the number of clusters of size k at time t.
- Introduce the random variables

$$w_{n,k}(t) := \frac{1}{n} \# \{ \text{clusters of size } k \text{ at time } t \}, \quad 1 \leq k \leq n.$$

Rather than working with these quantities directly, use the

$$G_n(x,t) = \sum_{k=1}^n x^k w_{n,k}(t), \qquad n \ge 1, x \in (0,1), t \ge 0$$

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 Rather than working with these quantities directly, use the empirical generating function

$$G_n(x,t) = \sum_{k=1}^n x^k w_{n,k}(t), \qquad n \ge 1, x \in (0,1), t \ge 0$$

#### Theorem 1

#### Theorem

Suppose that the coalescence rates  $\alpha: \mathbb{N} \to \mathbb{R}^+$  satisfy

$$\alpha(k) \le \exp(\gamma k \ln \ln(k)), \quad \forall k,$$

where  $\gamma < 1$  is an arbitrary constant. Let  $G : [0,1] \times \mathbb{R}^+ \to \mathbb{R}$  be the solution of the deterministic initial value problem

$$G(x,0) = x,$$

$$\frac{\partial G}{\partial t}(x,t) = \lambda(x - G(x,t)) + \sum_{k=0}^{\infty} \frac{\alpha(k)}{k!} \left( G(x,t)^k - kG(1,t)^{k-1} G(x,t) \right).$$

Then  $G_n(x,t)$  converges to G(x,t) in  $L^2$ , uniformly in x and t, as  $n \to \infty$ , that is

$$\sup_{x\in[0,1],t\geq0}\mathbb{E}\left[\left(G(x,t)-G_n(x,t)\right)^2\right]\to0,\quad \text{as}\quad n\to\infty.$$

• For  $f(x, \mathbf{w_n}) := \sum_{k=1}^n x^k w_{n,k}$ , we have

$$\mathcal{A}_{n}f(x, \mathbf{w}_{n}) = \lambda(x - f(x, \mathbf{w}_{n}))$$

$$+ \sum_{k=2}^{n} \frac{\alpha(k)}{k!} (f(x, \mathbf{w}_{n})^{k} - kf(1, \mathbf{w}_{n})^{k-1} f(x, \mathbf{w}_{n}))$$

$$+ \beta_{n}(x, \mathbf{w}_{n}),$$

where

$$\sup_{\mathbf{w}_n} |\beta_n(\mathbf{x}, \mathbf{w}_n)| \leq \frac{A}{n},$$

where A is a constant independent of n and x.

## Main technique in proof

- Look at the mean-field equations to "guess" the limiting behaviour of  $G_n(x,t)$  (equivalently consider the leading order terms of the generator).
- Apply Dynkin's formula, play with leading terms in generator

$$\begin{split} \mathbb{E}[(G(x,t) - G_n(x,t))^2] \\ &= \mathbb{E}\left[\int_0^t \left(\frac{\partial}{\partial s} + \mathcal{A}_n\right) [(G(x,s) - G_n(x,s))^2] ds\right], \end{split}$$

## Main technique in proof

- Look at the mean-field equations to "guess" the limiting behaviour of  $G_n(x,t)$  (equivalently consider the leading order terms of the generator).
- Apply Dynkin's formula, play with leading terms in generator and invoke Gronwall's Lemma:

$$\begin{split} \mathbb{E}[(G(x,t)-G_n(x,t))^2] \\ &= \mathbb{E}\left[\int_0^t \left(\frac{\partial}{\partial s} + \mathcal{A}_n\right) [(G(x,s)-G_n(x,s))^2] ds\right], \end{split}$$

- The next theorem deals with the stationary cluster size distribution.
- Let

$$p_{n,k}(t) := \frac{\#\{\text{clusters of size } k \text{ at time } t\}}{\#\{\text{clusters at time } t\}}, \quad 1 \leq k \leq n.$$

Define

$$p_k := \lim_{t \to \infty} \lim_{n \to \infty} p_{n,k}(t)$$
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as a distributional limit, which exists thanks to the previous theorem and that

$$\sum_{k=1}^{n} x^{k} p_{n,k}(t) = \frac{G_{n}(x,t)}{G_{n}(1,t)}, \qquad n \ge 1, x \in (0,1), t \ge 0.$$

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#### Theorem 2

#### Theorem

If  $\alpha$  satisfies

$$\alpha(k) \le \exp(\gamma k \ln \ln(k)), \quad \forall k$$

and m is the smallest integer such that  $\alpha(m) > 0$ , then the stationary cluster size distribution obeys

$$\lim_{\lambda \searrow 0} p_k = \begin{cases} \frac{1}{k} \left(\frac{m-1}{m}\right)^k \left(\frac{1}{m}\right)^{\frac{k-1}{m-1}} {m \choose \frac{k-1}{m-1}} \end{cases} \quad \text{if } m-1 \text{ divides } k-1 \\ 0 \qquad \qquad \text{otherwise} \end{cases}$$

and in particular, as  $k \to \infty$ 

$$\lim_{\lambda \searrow 0} p_k pprox egin{cases} k^{-3/2} & \textit{if } m-1 \textit{ divides } k-1 \\ 0 & \textit{otherwise}. \end{cases}$$

- Suppose we allow coalescence in groups of three or more but not pairs (m=3).
- In the large n and small  $\lambda$  limit we will see no clusters of even
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- This is a consequence of the weight of the tail of the cluster size distribution.
- The universal exponent 3/2 suggests a typical cluster size  $\sum_{1}^{n} k p_{k} \approx O(n^{1/2}) \Rightarrow \sharp \text{ clusters } \approx O(n^{1/2}).$ Coalescence of triples:  $\binom{n^{1/2}}{3} \times \alpha(3) n^{1-3} \approx O(n^{-1/2})$

Coalescence of quadruples:  $\binom{n^{1/2}}{4} \times \alpha(4) n^{1-4} \approx O(n^{-1})$ 

With 2/3 of blocks being singletons, this creates an imbalance with manifests in the disappearance of even sized blocks.

Some more strange results for exchangeable fragmentation-coalescence models<sup>2</sup>

<sup>&</sup>lt;sup>2</sup> Joint work with Steven Pagett, Tim Rogers and Jason Schweinsberg.



## Kingman *n*-coalescent

• The Kingman *n*-coalescent is an (exchangeable) coalescent process on the space of partitions of  $\{1, \dots, n\}$  denoted by

$$\Pi^{(n)}(t) = (\Pi_1^{(n)}(t), \cdots, \Pi_{N(t)}^{(n)}(t)), \qquad t \ge 0,$$

where N(t) is the number of blocks at time t and  $\Pi_i^{(n)}(t)$  is the elements of  $\{1, \dots, n\}$  that belong to the *i*-th block.

- Blocks merge in pairs, with a fixed rate c of any two blocks
- Both N(t),  $t \ge 0$ , is a Markov process and  $\Pi^{(n)}$  is a Markov
- The notion of the Kingman coalescent can be mathematically

$$\{\Pi(t): t \ge 0\} := \lim_{n \to \infty} \{\Pi^{(n)}(t): t \ge 0\}$$

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- Blocks merge in pairs, with a fixed rate c of any two blocks merging.
- Both N(t),  $t \ge 0$ , is a Markov process and  $\Pi^{(n)}$  is a Markov process.
- The notion of the Kingman coalescent can be mathematically extended in a consistent way to the space of partitions on  $\mathbb{N}$ . That is to say the pathwise limit

$$\{\Pi(t): t \ge 0\} := \lim_{n \to \infty} \{\Pi^{(n)}(t): t \ge 0\}$$

## Kingman coalescent

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- (Slighly) more precisely: if the initial configuration is the

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## Kingman coalescent

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- (Slighly) more precisely: if the initial configuration is the trivial partition

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(so that  $N(0) = \infty$ ) then  $N(t) < \infty$  almost surely, for all t > 0.

• In particular, the Markov Chain N(t) has an entrance law at  $+\infty$ .

- At rate  $\mu$ , each block in the system is shattered into singletons.
- When there are a finite number of blocks, each block must contain an infinite number of integers and hence when a block shatters, the system jumps back up to "infinity".
- If started with a finite number of blocks, the resulting process is still a Markov process on the space of partitions of  $\mathbb N$  until the arrival of the first fragmentation.
- Can process be "extended" to a Markov process on N∪ {+∞}? Can the process "come down from infinty"?
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• We can continue to use the same notation as before with

$$\Pi(t) = (\Pi_1(t), \cdots \Pi_{N(t)}), \qquad t \ge 0,$$

as a partitioned-valued process.

- A little thought (exchangeability!) shows that both N(t) and M(t) := 1/N(t),  $t \ge 0$ , are Markov process (with a possible absorbing state at  $+\infty$  resp. 0).
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### A remarkable phase transition

#### $\mathsf{Theorem}$

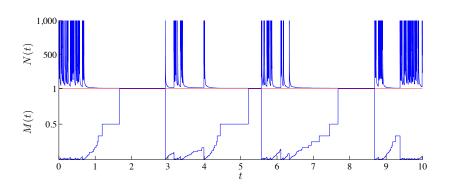
If  $\theta := 2\mu/c < 1$ , then M is a recurrent strong Markov process on  $\{1/n : n \in \mathbb{N}\} \cup \{0\}.$ 

(Comes down from infinity.)

If  $\theta := 2\mu/c \ge 1$ , then 0 is an absorbing state for M.

(Does not come down from infinity.)

## Coming down from infinity



- We can build an excursion theory for N (resp. M) at  $\infty$  (resp. 0). In particular there exists a local time L of N (resp. M) at  $\infty$  (resp. 0).
- Zero time at the boundary point:
- Aforesaid inverse local time has Laplace exponent

$$\Phi(q) = \frac{\Gamma(1-\theta)\Gamma(1-\alpha^+(q))\Gamma(1-\alpha^-(q))}{\Gamma(\alpha^+(q))\Gamma(\alpha^-(q))}, \qquad q \ge 0,$$

$$\alpha^{\pm}(q) = \frac{1-\theta}{2} \pm \frac{1}{2}\sqrt{(1+\theta)^2 - 8q/c}, \qquad q \ge 0.$$

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## Stationary distribution

#### **Theorem**

Let  $\theta:=2\mu/c<1$ , then M has stationary distribution given by the Beta-Geometric  $(1-\theta,\theta)$  distribution

$$\rho_{M}(1/k) = \frac{(1-\theta)}{\Gamma(\theta)} \frac{\Gamma(k-1+\theta)}{\Gamma(k+1)}, \quad k \in \mathbb{N}.$$

In particular  $\rho_M(0) = 0$ .

Thank you!